Improving long-term survival estimation through flexible models, combining evidence and accessible software

Chris Jackson MRC Biostatistics Unit, Cambridge, U.K.

UCL, 7 July 2015

"Improving long-term survival estimation through

- 1. flexible models
- 2. combining evidence
- 3. accessible software. . .

. . . but you can't currently have all three at the same time!"

"Improving long-term survival estimation through

- 1. flexible models
- 2. combining evidence
- 3. accessible software. . .

. . . but you can't currently have all three at the same time!"

Premise of the talk

Long-term survival modelling needs

- \blacktriangleright Sufficient long-term data:
	- \blacktriangleright individual data with long follow up, or
	- \blacktriangleright individual data with short follow up $+$ other data on the long-term period
- \triangleright Models flexible enough to capture the data
- \blacktriangleright Usable software, skills.

1. Using long-term population data to extrapolate RCT evidence over time (slides [6–](#page-6-0)[21\)](#page-28-0)

(Benaglia, Jackson & Sharples, Stat. Med (2015) 34(5):796–811)

flexible modelling \parallel combining evidence \parallel accessible software

- 2. Flexible parametric survival models for one individual dataset (slides [23](#page-34-0)[–37\)](#page-64-0)
	- \triangleright The flexsurv R package for parametric survival modelling

flexible modelling | combining evidence | accessible software

Part I

[Using long-term population data to](#page-5-0) [extrapolate short-term survival data](#page-5-0)

ICD (Implantable Cardioverter Defibrillators) compared to anti-arrhythmic drugs (AAD) for prevention of sudden cardiac death in people with cardiac arrhythmia.

Data:

- \blacktriangleright Individual data from cohort of 535 UK cardiac arrhythmia patients implanted with ICDs between 1991 and 2002.
- \triangleright Meta-analysis of three (non-UK) RCTs (published HRs).
	- \blacktriangleright Relatively short-term follow-up: approximately 75% people followed for less than 5 years, maximum 10 years

 \triangleright UK population mortality statistics by age, sex, cause of death. Estimate the survival curve over the lifetime of ICD and AAD patients in UK

Previous Work: central idea (Demiris & Sharples, Stat. Med. 2006)

Years

Use UK population data with same age/sex distribution to anchor the ICD population risk

Previous Work: central idea (Demiris & Sharples, Stat. Med. 2006)

Years

Use UK population data with same age/sex distribution to anchor the ICD population risk

Previous Work: key assumption

Log−Hazard Function

$h_{\mathsf{ICD}}(t) = \mathrm{e}^{\beta}h_{\mathsf{UK}}(t)$, for $t > 0$

Constant (multiplicative) hazard ratio between ICD and UK population

This seems a strong assumption:

- 1. ICD patients at greater risk of arrhythmia death
- 2. If proportion of deaths caused by arrhythmia changes over time, then extrapolating constant HR for all causes may be inaccurate

Previous Work: key assumption

Log−Hazard Function

Constant (multiplicative) hazard ratio between ICD and UK population

This seems a strong assumption:

- 1. ICD patients at greater risk of arrhythmia death
- 2. If proportion of deaths caused by arrhythmia changes over time, then extrapolating constant HR for all causes may be inaccurate

$$
h_{ICD}(t) = e^{\beta}h_{UK}(t)
$$
, for $t > 0$

Proportion of UK deaths which are due to arrhythmia

Proportion of Arrhythmic Deaths - UK Population 2002

Extrapolate constant cause-specific instead of all-cause hazard ratio

Simulation study: ignoring cause-specific nature of hazard gives bias in mean survival

- \triangleright particularly when hazard increases much quicker for other cause
- \triangleright See paper for full details (Benaglia, Jackson & Sharples, Stat. Med (2015) 34(5):796–811)

Application to ICD example. . .

Model to extrapolate survival for ICD patients

(not considering AAD control group, RCT data for the moment. . .)

General population data: cause of death $(k = \text{arrhythmic})$, non-arrhythmic) known.

Cause-specific survival is Weibull with hazard:

$$
h_{UK}^{(k)}(t) = \alpha_k \lambda_k t^{\alpha_k - 1}
$$

 \blacktriangleright ICD cohort: cause of death unknown Overall survival follows a polyhazard model (Louzada-Neto, Biometrics 1999):

$$
h_{\text{ICD}}(t) = h_{\text{ICD}}^{\text{arr}}(t) + h_{\text{ICD}}^{\text{other}}(t)
$$

 \triangleright t: minimum time to one of 2 possible causes of death

 \blacktriangleright Hazard is the sum of 2 cause-specific hazards

Cause-specific proportional hazards assumption

ICD cohort hazard is related to the general population hazard as:

$$
h_{ICD}(t) = h_{ICD}^{arr}(t) + h_{ICD}^{other}(t)
$$

= $e^{\beta} h_{UK}^{arr}(t) + h_{UK}^{other}(t)$
= $e^{\beta} \alpha_1 \lambda_1 t^{\alpha_1 - 1} + \alpha_2 \lambda_2 t^{\alpha_2 - 1}$ (poly-Weibull)

Arrhythmia hazard is proportional Other-cause hazard is identical to UK matched population.

- \triangleright Joint Bayesian model for ICD cohort $+$ UK population data
- Estimate joint posterior of parameters $\alpha_1, \alpha_2, \lambda_1, \lambda_2, \beta$ by MCMC (using WinBUGS).
- \triangleright WBDev add-on needed to implement the poly-Weibull distribution for the cohort data

Express beliefs on an intuitive scale — exact choice may make a difference for small populations Weibull rate λ :

- Age around 60 on study entry: cannot survive more than 60 additional years. Mean survival $\sim U(0, 60)$.
- ► $1/\lambda \sim U(0, 100)$, gives a mean $1/\lambda \Gamma(1 + 1/\alpha)$ of < 60 , for all plausible α .

Weibull shape α : controls hazard vs. time: $h(t)=\alpha\lambda(\lambda t)^{\alpha-1}$

- ► Hazard ratio for doubled time t is $2^{\alpha-1}$.
- \triangleright Prior mean of 1.5 for this, with 95% CI about (0.64, 100)
- \triangleright implies log(α) ∼ N(0.5, $\sigma = 0.78$)

Log HR β between ICD patients and general population: 95% CI for HR $(1/150,150) \rightarrow \beta \sim N(0,\sigma=2.5)$

Express beliefs on an intuitive scale — exact choice may make a difference for small populations Weibull rate λ :

- \triangleright Age around 60 on study entry: cannot survive more than 60 additional years. Mean survival $\sim U(0, 60)$.
- \blacktriangleright 1/ $\lambda \sim U(0, 100)$, gives a mean $1/\lambda \Gamma(1 + 1/\alpha)$ of < 60 , for all plausible α .

Weibull shape α : controls hazard vs. time: $h(t)=\alpha\lambda (\lambda t)^{\alpha-1}$

- ► Hazard ratio for doubled time t is $2^{\alpha-1}$.
- \triangleright Prior mean of 1.5 for this, with 95% CI about (0.64, 100)
- \triangleright implies log(α) ∼ N(0.5, $\sigma = 0.78$)

Log HR β between ICD patients and general population: 95% CI for HR $(1/150,150) \rightarrow \beta \sim N(0,\sigma=2.5)$

Express beliefs on an intuitive scale — exact choice may make a difference for small populations Weibull rate λ :

- \triangleright Age around 60 on study entry: cannot survive more than 60 additional years. Mean survival $\sim U(0, 60)$.
- \blacktriangleright 1/ $\lambda \sim U(0, 100)$, gives a mean $1/\lambda \Gamma(1 + 1/\alpha)$ of < 60 , for all plausible α .

Weibull shape α : controls hazard vs. time: $h(t)=\alpha\lambda (\lambda t)^{\alpha-1}$

- ► Hazard ratio for doubled time t is $2^{\alpha-1}$.
- \triangleright Prior mean of 1.5 for this, with 95% CI about (0.64, 100)
- \triangleright implies log(α) ∼ N(0.5, $\sigma = 0.78$)

Log HR β between ICD patients and general population: 95% CI for HR $(1/150,150) \rightarrow \beta \sim N(0, \sigma = 2.5)$

Extrapolating real ICD cohort data

- I Ignore the cause-specific hazard (Weibull) or account for it (poly-Weibull)
- \triangleright More bias for women when ignoring it
	- \blacktriangleright due to time-varying proportion of deaths due to arrhythmia.

Proportion of UK deaths which are due to arrhythmia

Proportion of Arrhythmic Deaths - UK Population 2002

Including an intervention effect from literature

Hazards for three groups under Poly-Weibull model:

$$
h_{\text{UK}}(t) = h_{\text{UK}}^{arr}(t) + h_{\text{UK}}^{other}(t)
$$

\n
$$
h_{\text{ICD}}(t) = e^{\beta} h_{\text{UK}}^{arr}(t) + h_{\text{UK}}^{other}(t)
$$

\n
$$
h_{\text{AAD}}(t) = e^{\gamma_s + \beta} h_{\text{UK}}^{arr}(t) + h_{\text{UK}}^{other}(t),
$$

Meta-analysis of ICD vs AAD trials, published HR for arrhythmia mortality, gives a prior for γ_a .

For the (probably biased) Weibull model we have:

$$
h_{\text{UK}}(t) = \mu_1 \alpha t^{\alpha - 1} = e^{\beta_0} \alpha t^{\alpha - 1}
$$

\n
$$
h_{\text{ICD}}(t) = e^{\beta_1} h_{\text{UK}}(t) = e^{\beta_0 + \beta_1} \alpha t^{\alpha - 1}
$$

\n
$$
h_{\text{AAD}}(t) = e^{\gamma} h_{\text{ICD}}(t) = e^{\beta_0 + \beta_1 + \gamma} \alpha t^{\alpha - 1},
$$

Prior for γ from published meta-analysis HR for all-cause mortality. Outcome of interest \rightarrow life years gained (LYG) by ICDs vs AADs.

Chris Jackson, MRC-BSU Cambridge [Improving long-term survival estimation](#page-0-0) 16/37

Including an intervention effect from literature

Hazards for three groups under Poly-Weibull model:

$$
h_{\text{UK}}(t) = h_{\text{UK}}^{arr}(t) + h_{\text{UK}}^{other}(t)
$$

\n
$$
h_{\text{ICD}}(t) = e^{\beta} h_{\text{UK}}^{arr}(t) + h_{\text{UK}}^{other}(t)
$$

\n
$$
h_{\text{AAD}}(t) = e^{\gamma_s + \beta} h_{\text{UK}}^{arr}(t) + h_{\text{UK}}^{other}(t),
$$

Meta-analysis of ICD vs AAD trials, published HR for arrhythmia mortality, gives a prior for γ_a .

For the (probably biased) Weibull model we have:

$$
h_{\text{UK}}(t) = \mu_1 \alpha t^{\alpha - 1} = e^{\beta_0} \alpha t^{\alpha - 1}
$$

\n
$$
h_{\text{ICD}}(t) = e^{\beta_1} h_{\text{UK}}(t) = e^{\beta_0 + \beta_1} \alpha t^{\alpha - 1}
$$

\n
$$
h_{\text{AAD}}(t) = e^{\gamma} h_{\text{ICD}}(t) = e^{\beta_0 + \beta_1 + \gamma} \alpha t^{\alpha - 1},
$$

Prior for γ from published meta-analysis HR for all-cause mortality. Outcome of interest \rightarrow life years gained (LYG) by ICDs vs AADs.

- \blacktriangleright ICD cohort extrapolated using population data
- \triangleright AAD survival generated with aid of meta-analysis.
- \blacktriangleright Life-years gained from ICD appears biased if use Weibull

Extrapolating incremental survival between interventions

- \blacktriangleright ICD cohort extrapolated using population data
- \triangleright AAD survival generated with aid of meta-analysis.
- \blacktriangleright Life-years gained from ICD appears biased if use Weibull
- \blacktriangleright Slightly more apparent bias for women

Chris Jackson, MRC-BSU Cambridge [Improving long-term survival estimation](#page-0-0) 1996 17/ 37

Proportion of UK deaths which are due to arrhythmia

Proportion of Arrhythmic Deaths - UK Population 2002

Causes of death may be recorded inconsistently between

- \triangleright meta-analysis of ICD vs drug trials "HR for arrhythmia deaths"
- \blacktriangleright population mortality data

Sensitivity analysis — assume 10%-20%

"arrhythmia"/"non-arrhythmia" deaths are misclassified.

- \triangleright e.g. if fewer deaths actually affected by treatment, expected survival gains from treatment lower
- \triangleright Still doesn't explain the discrepancy between models for men.

Causes of death may be recorded inconsistently between

- \triangleright meta-analysis of ICD vs drug trials "HR for arrhythmia deaths"
- \blacktriangleright population mortality data

Sensitivity analysis — assume 10%-20%

"arrhythmia"/"non-arrhythmia" deaths are misclassified.

- \triangleright e.g. if fewer deaths actually affected by treatment, expected survival gains from treatment lower
- \triangleright Still doesn't explain the discrepancy between models for men.

We assumed ICD patients had

 $\left\{ \text{arrhythmia hazard proportional}(\text{=greater}) \atop \text{other-cause hazard identical} \right\}$ to general population.

- \triangleright What if ICD patients at greater risk from some other causes (other heart disease), as well as arrhythmia?
- \triangleright May have led to biases in survival (underestimation of AAD-specific survival in poly-Weibull model. . .reasoning for this in paper)

- \triangleright Bayesian models useful for combining short-term RCT / cohort and longer-term survival data.
- \blacktriangleright Ignoring cause-specific hazard, thus misspecifying the underlying model, introduces bias in survival estimates.
	- \triangleright may underestimate or overestimate overall survival.
- \triangleright Bias can be alleviated by modelling cause-specific hazards
	- \triangleright but requires cause-specific survival data / treatment effects
	- \triangleright and information about which causes will be affected by disease status and / or treatment
- \triangleright Bias for treatment comparisons may be less if bias acts in the same way in all treatment groups.
- \triangleright Sensitivity analysis to model / data assumptions important
- \triangleright Requires model-specific BUGS code...

- \triangleright Bayesian models useful for combining short-term RCT / cohort and longer-term survival data.
- I Ignoring cause-specific hazard, thus misspecifying the underlying model, introduces bias in survival estimates.
	- \triangleright may underestimate or overestimate overall survival.
- \triangleright Bias can be alleviated by modelling cause-specific hazards
	- \triangleright but requires cause-specific survival data / treatment effects
	- \triangleright and information about which causes will be affected by disease status and / or treatment
- \triangleright Bias for treatment comparisons may be less if bias acts in the same way in all treatment groups.
- \triangleright Sensitivity analysis to model / data assumptions important
- \triangleright Requires model-specific BUGS code...

- \triangleright Bayesian models useful for combining short-term RCT / cohort and longer-term survival data.
- I Ignoring cause-specific hazard, thus misspecifying the underlying model, introduces bias in survival estimates.
	- \triangleright may underestimate or overestimate overall survival.
- \triangleright Bias can be alleviated by modelling cause-specific hazards
	- \triangleright but requires cause-specific survival data / treatment effects
	- \triangleright and information about which causes will be affected by disease status and / or treatment
- \triangleright Bias for treatment comparisons may be less if bias acts in the same way in all treatment groups.
- \triangleright Sensitivity analysis to model / data assumptions important
- \triangleright Requires model-specific BUGS code...

- \triangleright Bayesian models useful for combining short-term RCT / cohort and longer-term survival data.
- I Ignoring cause-specific hazard, thus misspecifying the underlying model, introduces bias in survival estimates.
	- \triangleright may underestimate or overestimate overall survival.
- \triangleright Bias can be alleviated by modelling cause-specific hazards
	- \triangleright but requires cause-specific survival data / treatment effects
	- \triangleright and information about which causes will be affected by disease status and / or treatment
- \triangleright Bias for treatment comparisons may be less if bias acts in the same way in all treatment groups.
- \triangleright Sensitivity analysis to model / data assumptions important
- \triangleright Requires model-specific BUGS code...

- \triangleright Bayesian models useful for combining short-term RCT / cohort and longer-term survival data.
- I Ignoring cause-specific hazard, thus misspecifying the underlying model, introduces bias in survival estimates.
	- \triangleright may underestimate or overestimate overall survival.
- \triangleright Bias can be alleviated by modelling cause-specific hazards
	- \triangleright but requires cause-specific survival data / treatment effects
	- \triangleright and information about which causes will be affected by disease status and / or treatment
- \triangleright Bias for treatment comparisons may be less if bias acts in the same way in all treatment groups.
- \triangleright Sensitivity analysis to model / data assumptions important
- \triangleright Requires model-specific BUGS code...

Part II

[Flexible parametric survival models and](#page-33-0) [software](#page-33-0)

- \triangleright 2-parameter Weibull model judged adequate in ICD application $-$ but this is not always the case.
- \triangleright 3–4 parameter models (generalised gamma / F) (see e.g. Jackson et al Int J Biostat) sometimes better

More flexible is a spline-based model (Royston & Parmar, Stat. Med. 2002)

- \triangleright can model both baseline hazard and non-proportional hazards between groups with any number of parameters
	- \triangleright can fit as well as needed
- \triangleright can be implemented in Stata (stpm2) and R (flexsurv)

Log cumulative hazard (log(-log survival))

 \triangleright Weibull: linear function of log time t

Spline: piecewise cubic function of $log(t)$

 \blacktriangleright Pieces separated by knots to span the data

0 knots: Weibull

n knots: $n + 2$ parameters

(German breast cancer data, see Sauerbrei & Royston,

J.Roy.Stat.Soc A 1999)

Log cumulative hazard (log(-log survival))

 \triangleright Weibull: linear function of log time t

Spline: piecewise cubic function of $log(t)$

- \blacktriangleright Pieces separated by knots to span the data
- 0 knots: Weibull
- n knots: $n + 2$ parameters

(German breast cancer data, see Sauerbrei & Royston,

J.Roy.Stat.Soc A 1999)

(German breast cancer data, see Sauerbrei & Royston,

J.Roy.Stat.Soc A 1999)

Log cumulative hazard (log(-log survival))

 \triangleright Weibull: linear function of log time t

Spline: piecewise cubic function of $log(t)$

- \blacktriangleright Pieces separated by knots to span the data
- 0 knots: Weibull
- n knots: $n + 2$ parameters

Log cumulative hazard (log(-log survival))

 \triangleright Weibull: linear function of log time t

Spline: piecewise cubic function of $log(t)$

- \blacktriangleright Pieces separated by knots to span the data
- 0 knots: Weibull
- n knots: $n + 2$ parameters

(German breast cancer data, see Sauerbrei & Royston,

J.Roy.Stat.Soc A 1999)

(German breast cancer data, see Sauerbrei & Royston,

J.Roy.Stat.Soc A 1999)

(German breast cancer data, see Sauerbrei & Royston,

J.Roy.Stat.Soc A 1999)

Log cumulative hazard (log(-log survival))

 \triangleright Weibull: linear function of log time t

Spline: piecewise cubic function of $log(t)$

 \blacktriangleright Pieces separated by knots to span the data

0 knots: Weibull

(German breast cancer data, see Sauerbrei & Royston,

J.Roy.Stat.Soc A 1999)

Log cumulative hazard (log(-log survival))

 \triangleright Weibull: linear function of log time t

Spline: piecewise cubic function of $log(t)$

 \blacktriangleright Pieces separated by knots to span the data

0 knots: Weibull

(German breast cancer data, see Sauerbrei & Royston,

J.Roy.Stat.Soc A 1999)

Log cumulative hazard (log(-log survival))

 \triangleright Weibull: linear function of log time t

Spline: piecewise cubic function of $log(t)$

 \blacktriangleright Pieces separated by knots to span the data

0 knots: Weibull

(German breast cancer data, see Sauerbrei & Royston,

J.Roy.Stat.Soc A 1999)

Log cumulative hazard (log(-log survival))

 \triangleright Weibull: linear function of log time t

Spline: piecewise cubic function of $log(t)$

 \blacktriangleright Pieces separated by knots to span the data

0 knots: Weibull

Log cumulative hazard in terms of $x = log(t)$, parameters γ

$$
log(-log(S(x,\gamma))) = \gamma_0 + \gamma_1 x
$$

for Weibull

Log cumulative hazard in terms of $x = log(t)$, parameters γ

$$
log(-log(S(x, \gamma))) = \gamma_0 + \gamma_1 x + \gamma_2 v_1(x) + \ldots + \gamma_{m+1} v_m(x)
$$

for spline with *m* knots. $v_i(x)$ is *j*th *basis* function

Log cumulative hazard in terms of $x = log(t)$, parameters γ

$$
log(-log(S(x, \gamma))) = \gamma_0 + \gamma_1 x + \gamma_2 v_1(x) + \ldots + \gamma_{m+1} v_m(x)
$$

for spline with m knots. $v_j(x)$ is jth basis function

$$
v_j(x) = (x - k_j)_+^3 - \lambda_j (x - k_{min})_+^3 - (1 - \lambda_j)(x - k_{max})_+^3, \qquad \lambda_j = \frac{k_{max} - k_j}{k_{max} - k_{min}}
$$

and $(x - a)_+ = max(0, x - a)$.
a cubic specially constructed to give a smooth function at the knots
 k_{min}, \dots, k_{max}

Log cumulative hazard defined by

$$
log(-log(S(x,\gamma))) = \gamma_0 + \gamma_1 x + \gamma_2 v_1(x) + \ldots + \gamma_{m+1} v_m(x)
$$

Covariates z can be placed on any parameter γ

- $\blacktriangleright \ \gamma_0 = \boldsymbol{\beta}^\top \mathsf{z}$ gives a proportional hazards model
	- \triangleright Sufficiently flexible splines give similar answers to Cox models
- \triangleright Modelling any other γ_r as linear in z gives non-proportional
	- \triangleright hazard ratio can be an arbitrarily flexible function of time

Log cumulative hazard defined by

$$
log(-log(S(x,\gamma))) = \gamma_0 + \gamma_1 x + \gamma_2 v_1(x) + \ldots + \gamma_{m+1} v_m(x)
$$

Covariates z can be placed on any parameter γ

- $\blacktriangleright \ \gamma_0 = \boldsymbol{\beta}^\top \mathsf{z}$ gives a proportional hazards model
	- \triangleright Sufficiently flexible splines give similar answers to Cox models
- \triangleright Modelling any other γ_r as linear in z gives non-proportional hazards:
	- \blacktriangleright hazard ratio can be an arbitrarily flexible function of time

- \blacktriangleright improves short-term fit
- \blacktriangleright changes extrapolation

- \blacktriangleright improves short-term fit
- \blacktriangleright changes extrapolation

Interconnected issues:

- \blacktriangleright Hard to combine with evidence synthesis...
- \triangleright ...No easy Bayesian implementation (tricky in BUGS)...
- \blacktriangleright ... Priors for / interpretation of spline coefficients γ .

flexsurv R package

flexsurv (Jackson, CRAN (2011–2015))

- \triangleright Fit any parametric survival model to individual data by maximum likelihood
	- \triangleright right/interval censoring, left-truncation, multi-state models
- \triangleright Standard models built in, plus spline, generalized gamma/F
- \triangleright Can program your own novel parametric distributions, given the hazard or probability density function
- \triangleright Any parameter of any distribution can depend on covariates
	- \triangleright proportional hazards / accelerated failure time typical, but not necessary
- \triangleright Customisable outputs / summary presentation

Similar Stata packages

- \triangleright stpm2 (Royston & Lambert) for spline model
- \triangleright stgenreg (Crowther & Lambert) for user-defined parametric survival models

flexsurv R package

flexsurv (Jackson, CRAN (2011–2015))

- \triangleright Fit any parametric survival model to individual data by maximum likelihood
	- \triangleright right/interval censoring, left-truncation, multi-state models
- \triangleright Standard models built in, plus spline, generalized gamma/F
- \triangleright Can program your own novel parametric distributions, given the hazard or probability density function
- \triangleright Any parameter of any distribution can depend on covariates
	- \triangleright proportional hazards / accelerated failure time typical, but not necessary
- \triangleright Customisable outputs / summary presentation

Similar Stata packages

- \triangleright stpm2 (Royston & Lambert) for spline model
- \triangleright stgenreg (Crowther & Lambert) for user-defined parametric survival models

Fitting models, and default output

Generalized gamma 3-parameter model. Two out of the three parameters depend on a covariate (prognostic group) fs2 <- flexsurvreg(Surv(recyrs, censrec) ~ group + sigma(group), data=bc, dist="gengamma")

fs2

```
flexsurvreg(formula = Surv(recyrs, censrec) \tilde{z} group + sigma(group),
    data = bc, dist = "gengamma")
```


```
ATC = 1586.434
```
Fitting models, and default output

Generalized gamma 3-parameter model. Two out of the three parameters depend on a covariate (prognostic group) fs2 <- flexsurvreg(Surv(recyrs, censrec) ~ group + sigma(group),

```
data=bc, dist="gengamma")
```
fs2

```
Ca11:flexsurvreg(formula = Surv(recyrs, censrec) ~ group + sigma(group),
    data = bc, dist = "gengamma")
```
Estimates:


```
N = 686, Events: 299, Censored: 387
Total time at risk: 2113.425
Log-likelihood = -786.2172, df = 7ATC = 1586.434
```
Fitting models, plot output

 $plot(fs2, xlim=c(0, 10))$ $text(x=c(2, 3.5, 4), y=c(0.4, 0.55, 0.7),$ c("Poor","Medium","Good"))

Summarise user-defined functions of the parameters

e.g. restricted mean survival for a model with parameters α

$$
E(T|T
$$


```
mean.gengamma <- function(mu, sigma, Q,
                           horizon=10, \ldots}{
    surv \leq function(t, ...) {
         1 - pgengamma(q=t, mu=mu,
                        sigma=sigma, Q=Q, ...)
    }
    integrate(surv, 0, horizon, ...)$value
}
summary(fs2, newdata=list(group="Good"),
        t=1, fn=mean.gengamma)
```
Summarise user-defined functions of the parameters

e.g. restricted mean survival for a model with parameters α

$$
E(T|T
$$


```
mean.gengamma <- function(mu, sigma, Q,
                         horizon=10, \ldots}{
   surv \leq function(t, ...) {
        1 - pgengamma(q=t, mu=mu,
                      sigma=sigma, Q=Q, ...)
   }
   integrate(surv, 0, horizon, ...)$value
}
summary(fs2, newdata=list(group="Good"),
       t=1, fn=mean.gengamma)
group=Good
 time est lcl ucl
1 1 7.36671 6.725084 7.85355
```
- \triangleright Similar facility for people to fit their own parametric models
	- \triangleright given definitions of the hazard or probability density as R functions
- \triangleright User guide and Supplementary examples manuals give lots of worked examples
	- \triangleright User guide to be published in Journal of Statistical Software
- \triangleright Suggestions for additions welcome
	- \triangleright ideally build a user-contributed library of examples
- \triangleright Restricted to parametric modelling of one individual-level dataset
- \triangleright No random effects / frailty, though could be extended in theory

For synthesis of multiple related datasets, Bayesian probabilistic modelling languages (BUGS / JAGS / Stan) preferred

 \triangleright though higher barrier to entry at the moment

Modelling long term survival may need

- \blacktriangleright a combination of data sources...
	- \triangleright but modelling/software tools are tricky
	- \triangleright can be hard work to assess assumptions
- \blacktriangleright flexible models
	- \triangleright hard to combine with synthesis of multiple datasets
- \blacktriangleright accessible software

Modelling long term survival may need

- \blacktriangleright a combination of data sources...
	- \triangleright but modelling/software tools are tricky
	- \triangleright can be hard work to assess assumptions
- \blacktriangleright flexible models
	- \triangleright hard to combine with synthesis of multiple datasets
- \blacktriangleright accessible software

Modelling long term survival may need

- \blacktriangleright a combination of data sources...
	- \triangleright but modelling/software tools are tricky
	- \triangleright can be hard work to assess assumptions
- \blacktriangleright flexible models
	- \blacktriangleright hard to combine with synthesis of multiple datasets
- \blacktriangleright accessible software